Expectation of a binomial random variable with parameters p and n is pn. So, $E(Y|X = n) = \frac{n}{2}$ since $p = \frac{1}{2}$. Using law of total probability: $E(Y) = \sum_{n=1}^{6} E(Y|X = n)P(X = n) = \frac{1}{6} \sum_{n=1}^{6} \frac{n}{2} = \frac{7}{4}$ Expectation of a binomial random variable with parameters p and n is pn. So, $E(Y|X = n) = \frac{n}{2}$ since $p = \frac{1}{2}$. Using law of total probability: $E(Y) = \sum_{n=1}^{6} E(Y|X = n)P(X = n) = \frac{1}{6} \sum_{n=1}^{6} \frac{n}{2} = \frac{7}{4}$ Similarly,

$$E(XY) = \sum_{n=1}^{6} E(XY|X = n)P(X = n)$$

= $\sum_{n=1}^{6} E(nY|X = n)P(X = n)$
= $\frac{1}{6} \sum_{n=1}^{6} nE(Y|X = n)$
= $\frac{1}{6} \sum_{n=1}^{6} n\frac{n}{2} = \frac{91}{12}$

Three prisoners A, B, C are to be executed. The queen announces she will pardon one of them. She randomly picks one and tells the jailer but asks her not to reveal the name.

A sees the jailer and asks her who has been pardoned. She refuses to tell. A then asks who of B or C will be executed. The jailer thinks a bit and says 'B is being executed'.

Has the jailer's remark changed the probability of A being pardoned? Here is the reasoning of each party:

Jailer: The probability of any prisoner being pardoned is $\frac{1}{3}$. One of B or C must be executed so revealing that B is being executed did not change A's probability of being executed.

A: Since B will be executed only A or C can be pardoned so A's probability of being pardoned is $\frac{1}{2}$.

Who is right? I

 $=\frac{1}{2}$

Let A, B, C be the events that A, B, C respectively were pardoned. Let J be the event jailer says 'B is being executed'. We have to find $P(A|J) = \frac{P(A \cap J)}{P(J)}$ What happened is summarized below: Pardoned Jailer says 'B being executed' | 'C being executed' - both equiprobable i.e. $\frac{1}{2}$ А B 'C being executed' С 'B being executed' P(J) = P(jailer says 'B being executed') $P(J \cap A \text{ pardoned}) + P(J \cap C \text{ pardoned}) + P(J \cap B \text{ pardoned})$ $=\frac{1}{2} \times \frac{1}{3} + \frac{1}{3} \times 1 + \frac{1}{3} \times 0$

So,
$$P(A|J) = \frac{P(A \cap J)}{P(J)} = \frac{\frac{1}{6}}{\frac{1}{2}} = \frac{1}{3}$$

A wrongly interprets *J* as *B^c* giving $P(A|B^c) = \frac{P(A \cap B^c)}{P(B^c)} = \frac{\frac{1}{3}}{\frac{2}{3}} = \frac{1}{2}$

- In a coin tossing game you pay your friend Rs.10/- if your call is wrong and he pays you Rs.10/- if your call is correct.
- You lose Rs.100/- in 10 successive tosses where you called Heads but the coin turned up tails.
- Assuming the coin is fair what should you call on the eleventh toss?

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Since the coin is fair you can call either Heads or Tails it does not matter.

- There are 3 identical boxes (say 1, 2, 3) two of which contain Rs.10/- and the third has Rs.1000/-. The choice of boxes for putting the amounts is random. Also, the experimenter is told which box contains what amount.
- 2 You, the player, choose a particular box.
- 3 The experimenter chooses a second box and shows that it contains Rs.10/-. And gives you the option to switch to the remaining box, if you wish.
- 4 Should you switch?

Assume you choose Box 1. Let us analyse outcomes

Box 1	Box 2	Box 3	Stay	Switch
10	10	1000	10	1000
10	1000	10	10	1000
1000	10	10	1000	10

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Box 1	Box 2	Box 3	Stay	Switch
10	10	1000	10	1000
10	1000	10	10	1000
1000	10	10	1000	10

So, better to switch. Probability of winning Rs.1000/- is $\frac{2}{3}$ while staying has a probability of $\frac{1}{3}$.

Probability can play tricks with intuition

- Simpson's paradox. (Already discussed earlier)
- Gambler's fallacy.
- Monty Hall problem.

Examples: 4



Figure: The number of • is same on opposite sides

The game has two players You and I:

The players select one of the dice and roll. The winner is the one who rolls the higher number. Assume You select the die first.

Analysis with (die A, die B):

 $\mathcal{S}(S) = \{(2,1), (2,5), (2,9), (6,1), (6,5), (6,9), (7,1), (7,5), (7,9)\}$

The probability of each outcome is $\frac{1}{9}$. Probability Die A wins is $\frac{5}{9}$ and B wins is $\frac{4}{9}$. Repeating such an analysis with (die A, die C) and (die B, die C) gives the following table:

(die 1, die 2)	Prob. die 1 wins	Prob. die 2 wins
(A,B)	5/9	4/9
(A,C)	4/9	5/9
(B,C)	5/9	4/9

The player choosing first is always worse off since $A \succ B \succ C \succ A$ (where \succ is better than or higher probability than relation). Note that the \succ relation is not transitive.

Analysis: each die A. B rolled twice and summed. Let A, B be the random variables for the sums for A, B respectively.

 $B \in \{2, 6, 6, 10, 10, 10, 14, 14, 18\}$ and $A \in \{4, 8, 8, 9, 9, 12, 13, 13, 14\}$. We have 81 possible pairs as the outcome of the game. A's win probability is ³⁷/₈₁, B's is ⁴²/₈₁ and there is a tie with probability ²/₈₁.

A similar analysis for the others yields the following table:

(die 1, die 2)	Prob. die 1 wins	Prob. die 2 wins	Draw
(A,B)	37/81	42/81	2/81
(A,C)	44/81	33/81	4/81
(B,C)	37/81	42/81	2/81

In this case we have: $A \prec B \prec C \prec A$. Choosing first is still bad but the relations have reversed.

Probabilities do not follow intuitive rules.

Theorem 1 (Weak law of large numbers)

Sample average converges in probability towards the expected value as sample size $n \to \infty$. $\lim_{n \to \infty} P(|\bar{X}_n - \mu| > \epsilon) = 0$

Theorem 2 (Strong law of large numbers)

Sample average converges almost surely to the expected value as sample size $n \to \infty$. $P(\lim_{n \to \infty} \bar{X}_n = \mu) = 1$